

Biennial Convention 2007

# Adventures in Risk

23-26 September 2007 • Christchurch, New Zealand



Institute of Actuaries of Australia



# Translation invariant and positive homogeneous risk measures and portfolio management

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-It is wide-spread opinion  
the use of any **positive homogeneous  
and translation invariant** risk measure  
reduces to the same portfolio  
management as mean-variance risk  
measure does, which is equivalent to  
Markowitz portfolio

-We show that generally speaking  
it is not true, because these measures  
reduce to a **different** optimization problem.

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-We provide the explicit closed form solution of this problem

-The results will be demonstrated with the data of stocks from NASDAQ/Computer.



# Translation invariant and positive homogeneous risk measures

$$1) \quad \rho(X + \alpha) = \rho(X) + \alpha, \quad \alpha - \text{const}$$

$$2) \quad \rho(cX) = c\rho(X), \quad c > 0 - \text{const}$$

# Examples

## BASEL II

VaR

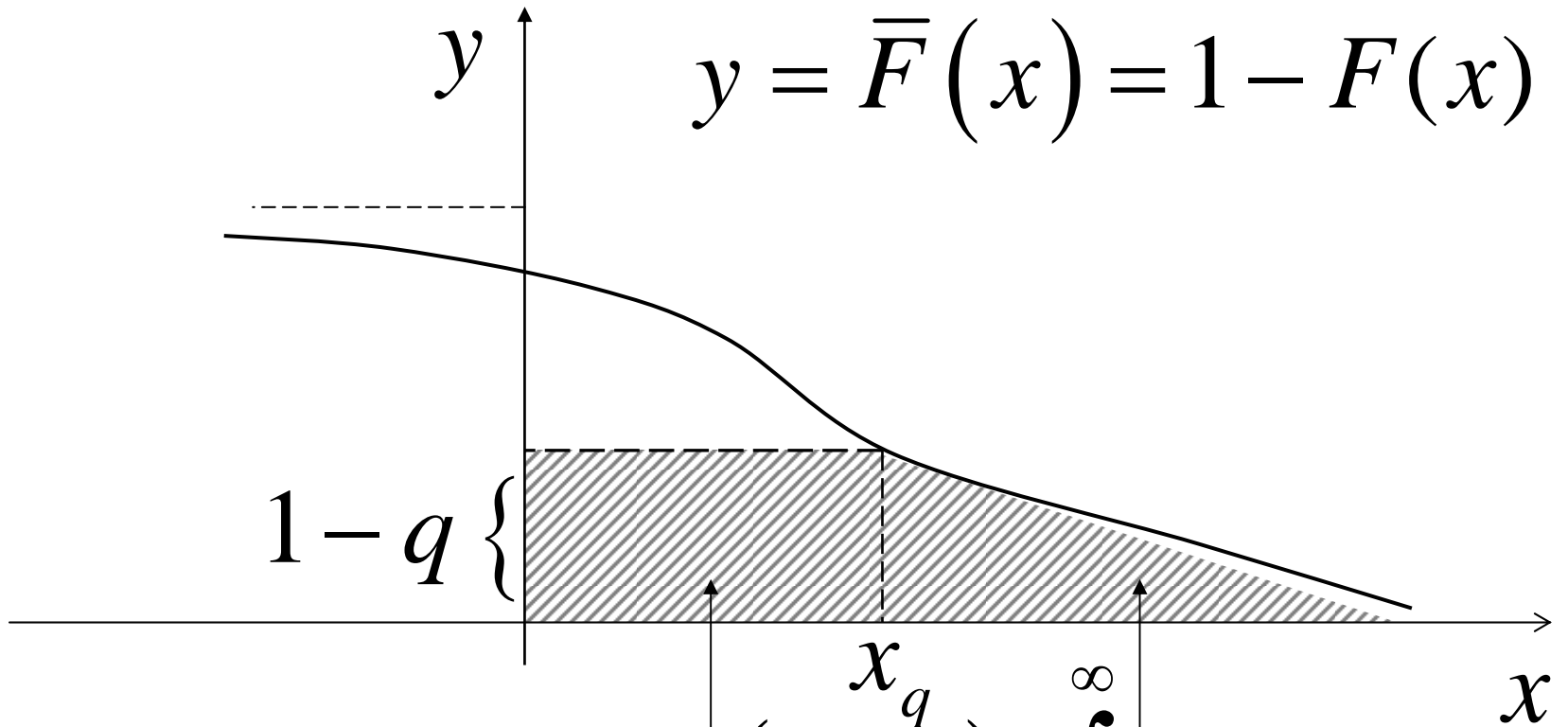
$$\begin{aligned}\rho(X) &= VaR_q(X) \\ &= \inf\{x \mid F_X(x) \geq q\}\end{aligned}$$

Expected Short Fall

$$\begin{aligned}\rho(X) &= ES_q(X) \\ &= E(X \mid X > VaR_q(X))\end{aligned}$$

Tail VaR, Tail conditional expectation (TCE)

<sup>5</sup>Conditional VaR (CVaR)



$$y = \bar{F}(x) = 1 - F(x)$$

$1 - q$  {

$x_q$

$$VaR_q(X) (1 - q) \int_{x_q}^{\infty} \bar{F}(x) dx$$

$$(1 - q) ES_q(X) = x_q (1 - q) + \int_{x_q}^{\infty} \bar{F}(x) dx$$





Panjer and Jing (2001) Hurlimann (2001)

L. and Valdez (2003, 2005)

## STD-premium

$$\rho(X) = E(X) + \lambda STD(X)$$

## Distorted risk measures

$$\rho(X) = \int_0^\infty g(\bar{F}(x)) dx,$$

Denneberg (1994) and Wang (1996)

## Coherent risk measures

Artzner, Delbean, Eber, and Heath (1999)



# *Tail SDP* Furman and Landsman (2006)

$$\rho(X) = E(X \mid X > VaR_q(X)) + \alpha \sqrt{V(X \mid X > VaR_q(X))}$$

Tail VaR

Tail Variance

Tail Expectation



$$ES_q(X)$$

$$= \arg \inf_c E((X - c)^2 \mid X > VaR_q(X))$$

$$V(X \mid X > VaR_q(X))$$

$$= \inf_c E((X - c)^2 \mid X > VaR_q(X))$$



## Portfolio management

$$X = (X_1, \dots, X_n)^T \square N_n(\mu, \Sigma)$$

$$P = \sum_{j=1}^n x_j X_j, \quad \sum_{j=1}^n x_j = 1$$

$$\rho(P) \rightarrow \inf$$

$$\rho(P) = \rho(\mathbf{x}^T X) = \rho(\mathbf{x}^T X - \mu^T \mathbf{x} + \mu^T \mathbf{x})$$

$$= \rho(\mathbf{x}^T X - \mu^T \mathbf{x}) + \mu^T \mathbf{x}$$

$$= \mu^T \mathbf{x} + \rho \left( \frac{\mathbf{x}^T X - \mu^T \mathbf{x}}{\sqrt{\text{VaR}(\mathbf{x}^T X)}} \sqrt{\text{VaR}(\mathbf{x}^T X)} \right)$$

$$Z \square N(0,1) = \mu^T \mathbf{x} + \rho \left( \frac{\mathbf{x}^T X - \mu^T \mathbf{x}}{\sqrt{\text{VaR}(\mathbf{x}^T X)}} \right) \sqrt{\text{VaR}(\mathbf{x}^T X)}$$

$$= \mu^T \mathbf{x} + \rho(Z) \sqrt{\mathbf{x}^T \Sigma \mathbf{x}} \rightarrow \inf$$

Translation invariant and positive homogeneous risk measure reduces to minimization

of linear plus **root** of quadratic functional

$$f(\mathbf{x}) = \mu^T \mathbf{x} + \lambda \sqrt{\mathbf{x}^T \Sigma \mathbf{x}} \rightarrow \inf$$

$$\mathbf{1}^T \mathbf{x} = 1$$

$$Z \square N(0, 1)$$

$$\lambda = Z_q \Leftrightarrow \inf VaR_q(P).$$

$$\lambda = ES_q(Z) \Leftrightarrow \inf ES_q(P).$$

$\rho(Z)$  – Any translation invariant and positive homogeneous risk measure

# Classical mean - variance optimization

$$q(\mathbf{x}) = E(P) + \lambda \text{Var}(P) = \boldsymbol{\mu}^T \mathbf{x} + \lambda \mathbf{x}^T \boldsymbol{\Sigma} \mathbf{x} \rightarrow \inf$$

linear + quadratic functional

subject to

$$\mathbf{1}^T \mathbf{x} = 1$$

Exact solution (see Boyle et.al (1998))

$$\mathbf{x}^* = \frac{\boldsymbol{\Sigma}^{-1} \mathbf{1}}{\mathbf{1}^T \boldsymbol{\Sigma}^{-1} \mathbf{1}} + \frac{1}{2\lambda} \left( \frac{\boldsymbol{\mu}^T \boldsymbol{\Sigma}^{-1} \mathbf{1}}{\mathbf{1}^T \boldsymbol{\Sigma}^{-1} \mathbf{1}} \boldsymbol{\Sigma}^{-1} \mathbf{1} - \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu} \right)$$

# Minimization of linear plus root of quadratic functional

We found the explicit closed form solution for minimization problem under system of equality constraints (L(2007))

$$\begin{cases} f(\mathbf{x}) = \mu^T \mathbf{x} + \lambda \sqrt{\mathbf{x}^T \Sigma \mathbf{x}} \rightarrow \inf \\ B\mathbf{x} = \mathbf{c} \end{cases}$$

$B = \text{matrix } m \times n, \quad \mathbf{c} = \text{vector}$

The special case for only one constraint  
 $m=1$

$$\begin{cases} f(\mathbf{x}) = \mu^T \mathbf{x} + \lambda \sqrt{\mathbf{x}^T \Sigma \mathbf{x}} \rightarrow \inf \\ \mathbf{1}^T \mathbf{x} = 1 \end{cases}$$

$$\Sigma = \begin{pmatrix} \Sigma_{11} & \sigma \\ \sigma^T & \sigma_{nn} \end{pmatrix},$$

$$\Sigma_{11} - (n-1) \times (n-1)$$



Define  $(n-1) \times (n-1)$  matrix

$$Q = \Sigma_{11} - \mathbf{1}_1 \sigma^T - \sigma \mathbf{1}_1^T + \sigma_{nn} \mathbf{1}_1 \mathbf{1}_1^T$$

**Lemma**  $\Sigma > 0 \Rightarrow Q > 0$ .

$$\Delta = (\mu_n - \mu_1, \dots, \mu_n - \mu_{n-1})^T.$$

**Theorem 1. L (2007a).** If

$$\lambda > \sqrt{\Delta^T Q^{-1} \Delta},$$



the minimization problem has the finite exact solution

$$\mathbf{x}^* = \frac{\Sigma^{-1}\mathbf{1}}{\mathbf{1}^T \Sigma^{-1}\mathbf{1}} + \frac{1}{\sqrt{(\lambda^2 - \Delta^T Q^{-1}\Delta)(\mathbf{1}^T \Sigma^{-1}\mathbf{1})}} (\Delta^T Q^{-1}, -\mathbf{1}_1^T Q^{-1}\Delta)^T$$



### 3. Value-at-Risk and Expected Short Fall

$$\lambda = ES_q(Z) = VaR_q(Z)$$

$$+ \frac{1}{1-q} \int_{VaR_q(Z)}^{\infty} (1-F(x))dx \geq VaR_q(Z) = Z_q.$$

# Theorem 2

If VaR-optimal portfolio solution is finite,  
the ES-optimal portfolio solution  
is automatically finite

Consider a portfolio of 10 stocks  
from NASDAQ/Computers

Table 1: Portfolio mean return

Stock	ADOBE	Compuware	NVIDIA	Staples	VeriSign
Mean	0.0061	-0.0081	-0.0096	0.0058	0.0064
Stock	Sandisk	Microsoft	Citrix	Intuit	Symantec
Mean	-0.0198	0.0002	-0.0038	-0.0041	0.0061



Table 2: Portfolio covariance return

	ADOBE	Compuware	NVIDIA	Staples	VeriSign
ADOBE	0.006102	0.001173	0.000118	0.000513	0.000121
Compuware	0.001173	0.003310	0.001047	0.000498	0.000847
NVIDIA	0.000118	0.001047	0.002145	0.000122	0.000772
Staples	0.000513	0.000498	0.000122	0.002940	-0.000547
VeriSign	0.000121	0.000847	0.000772	-0.000547	0.003486
	Sandisk	Microsoft	Citrix	Intuit	Symantec
Sandisk	0.004013	-0.000033	0.000844	0.000131	0.000083
Microsoft	-0.000033	0.000485	0.000220	0.000167	0.000062
Citrix	0.000844	0.000220	0.001365	0.000397	0.000445
Intuit	0.000131	0.000167	0.000397	0.000876	0.000027
Symantec	0.000083	0.000062	0.000445	0.000027	0.002542

# The loss on the portfolio

$$L = -R = -\sum_{j=1}^n x_j X_j,$$

$$VaR_q(L) = -\mu^T \mathbf{x} + z_q \sqrt{\mathbf{x}^T \Sigma \mathbf{x}} \xrightarrow{\mathbf{x}^T \mathbf{1}=1} \inf$$

$$ES_q(L) = -\mu^T \mathbf{x} + \lambda_q \sqrt{\mathbf{x}^T \Sigma \mathbf{x}} \xrightarrow{\mathbf{x}^T \mathbf{1}=1} \inf$$

From Theorem 1: Lower bound for solution

$$B = \sqrt{\Delta^T Q^{-1} \Delta} = 0.2142$$

$$X \square N_n(\mu, \Sigma)$$

VaR optimal portfolio finite  $q > 0.585$

ES optimal portfolio finite  $q > 0.112$   
 $q = 0.8$

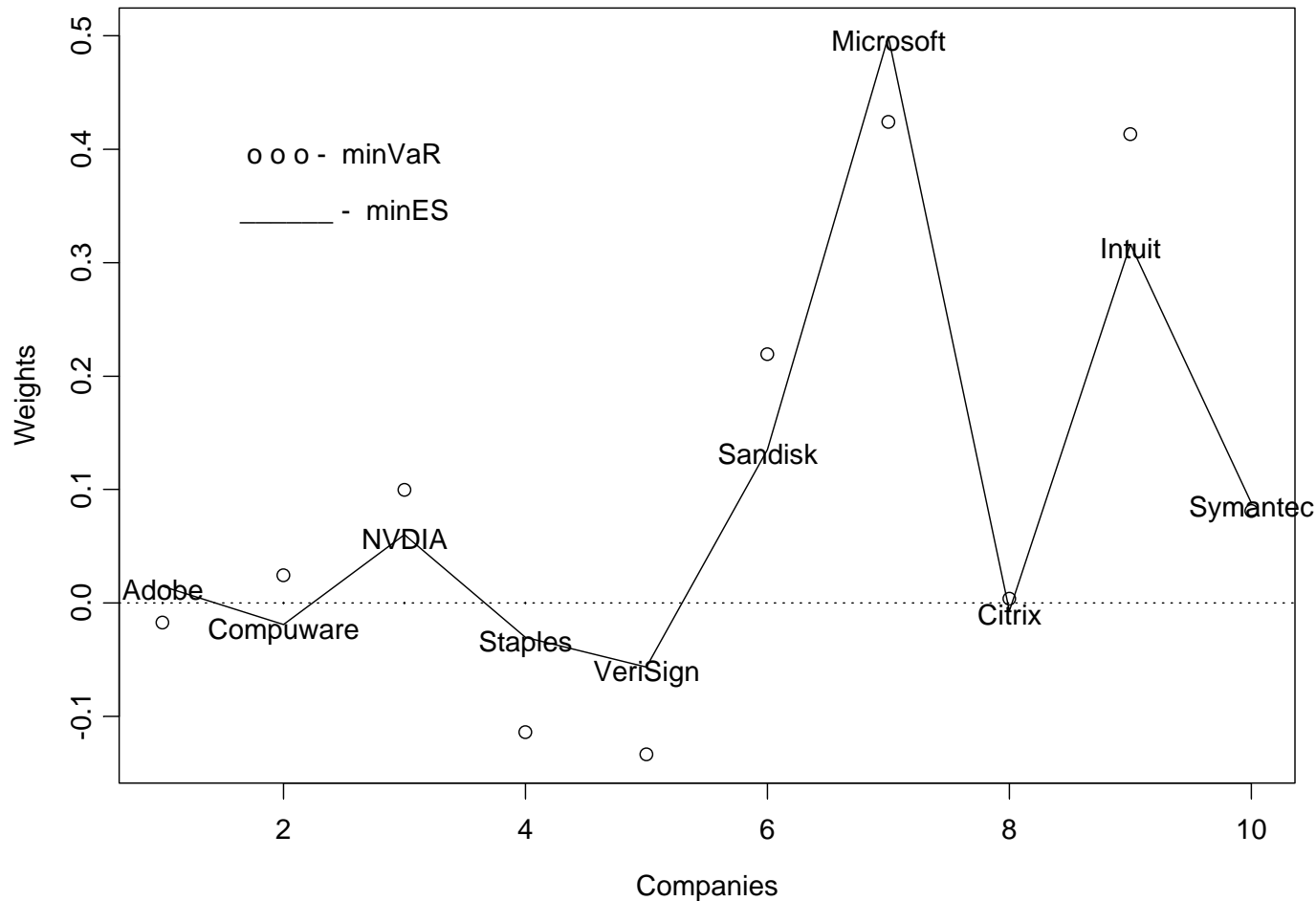
Table 3: Optimal portfolio

	ADOBE	Compuware	NVIDIA	Staples	VeriSign
minVaR	-0.0173	-0.024	0.0996	-0.1141	-0.1337
minTCE	0.0148	-0.0191	0.0601	-0.0305	-0.0567

	Sandisk	Microsoft	Citrix	Intuit	Symantec
minVaR	0.2192	0.4242	-0.0037	0.4133	0.081
minTCE	0.135	0.4993	-0.0071	0.3161	0.088

Table 4: Main characteristics for VaR and ES optimal portfolios

	VaR	TCE	Variance	Mean
minVaR	0.01109	0.0236	0.00052	0.00827
minTCE	0.01162	0.022	0.00035	0.0042



# Closeness between VaR and ES optimal portfolios



Distance between portfolios

$$d = \sqrt{(\arg \min_{\mathbf{1}^T \mathbf{x}} VaR_q(L) - \arg \min_{\mathbf{1}^T \mathbf{x}} ES_q(L))^T (\arg \min_{\mathbf{1}^T \mathbf{x}} VaR_q(L) - \arg \min_{\mathbf{1}^T \mathbf{x}} ES_q(L))}$$

Denote

$$C = (\mathbf{1}^T \Sigma^{-1} \mathbf{1})^{-1/2} \sqrt{\Delta^T Q^{-2} \Delta + (\mathbf{1}_1^T Q^{-1} \Delta)^2}$$





## Theorem 3

$$d = [(z_q^2 - B^2)^{-1/2} - (\lambda_q^2 - B^2)^{-1/2}]C \rightarrow 0,$$

$q \rightarrow 1$

Boundary

# Elliptical family

$$\mathbf{X} \square E_n(\boldsymbol{\mu}, \boldsymbol{\Sigma}, g_n).$$

$$f_{\mathbf{X}}(\mathbf{x}) = \frac{c_n}{\sqrt{|\boldsymbol{\Sigma}|}} g_n \left[ \frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right],$$

density generator

$$g_n(u) = \exp(-u) \square N_n(\boldsymbol{\mu}, \boldsymbol{\Sigma})$$



## Property

$$BX + c \sim E_m(B\mu + c, B\Sigma B^T, g_m)$$

guarantees

$$\rho(\mathbf{x}^T X) = \mu^T \mathbf{x} + \lambda \sqrt{\mathbf{x}^T \Sigma \mathbf{x}}$$

$$\lambda = \rho(Z), \quad Z \sim E_1(0, 1, g_1)$$

$$g_n(u) = \left(1 + u/k_{n,p}\right)^{-p} \square t_n(\boldsymbol{\mu}, \boldsymbol{\Sigma}, p)$$

## Multivariate Generalized t-distribution (GST)

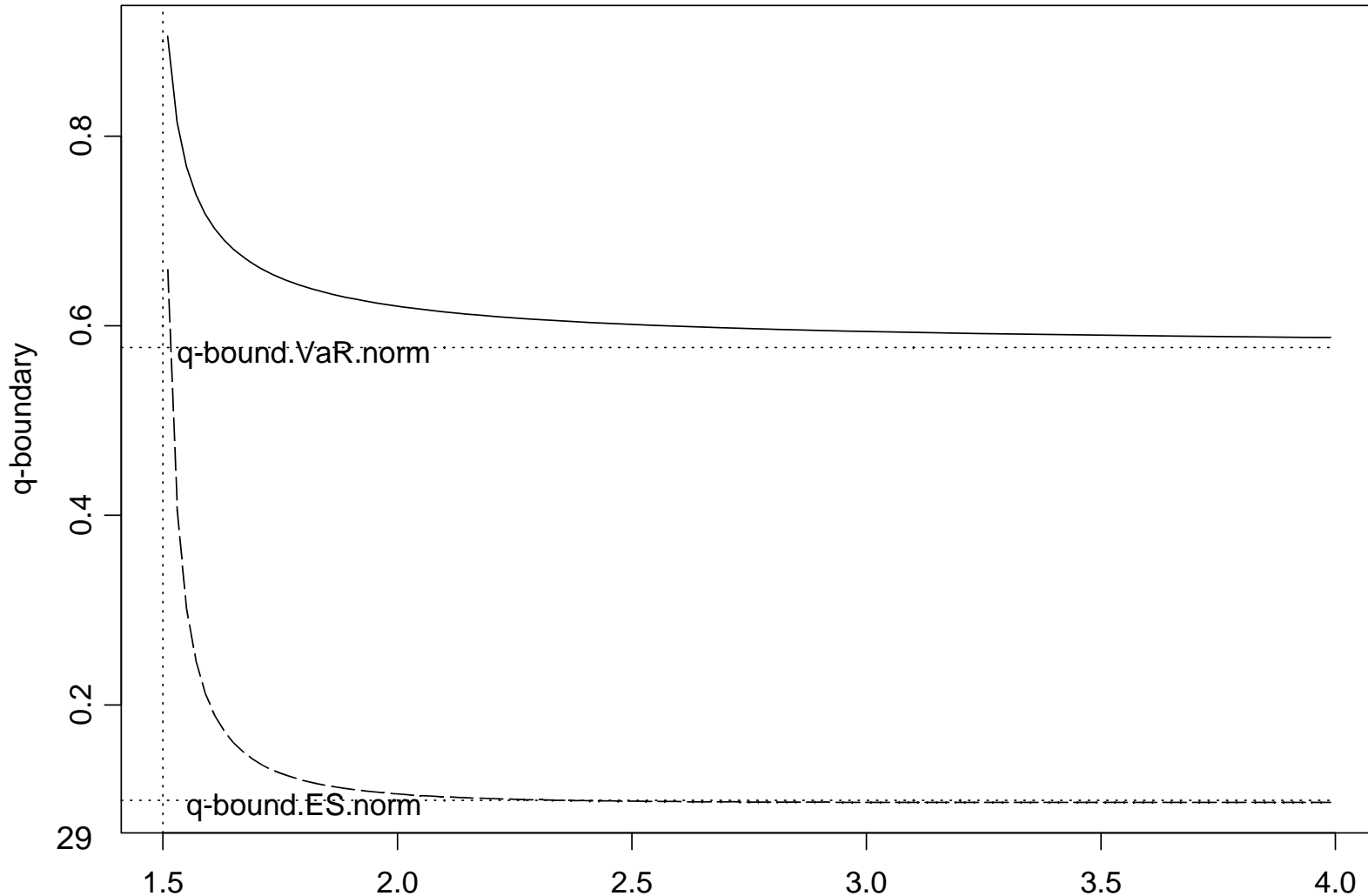
$$\begin{cases} f(\mathbf{x}) = -\boldsymbol{\mu}^T \mathbf{x} + \lambda \sqrt{\mathbf{x}^T \boldsymbol{\Sigma} \mathbf{x}} \rightarrow \inf \\ \mathbf{1}^T \mathbf{x} = 1 \end{cases}$$

$$T \square t_1(0, 1, p_1)$$

$$T_q \Rightarrow VaR - optimal$$

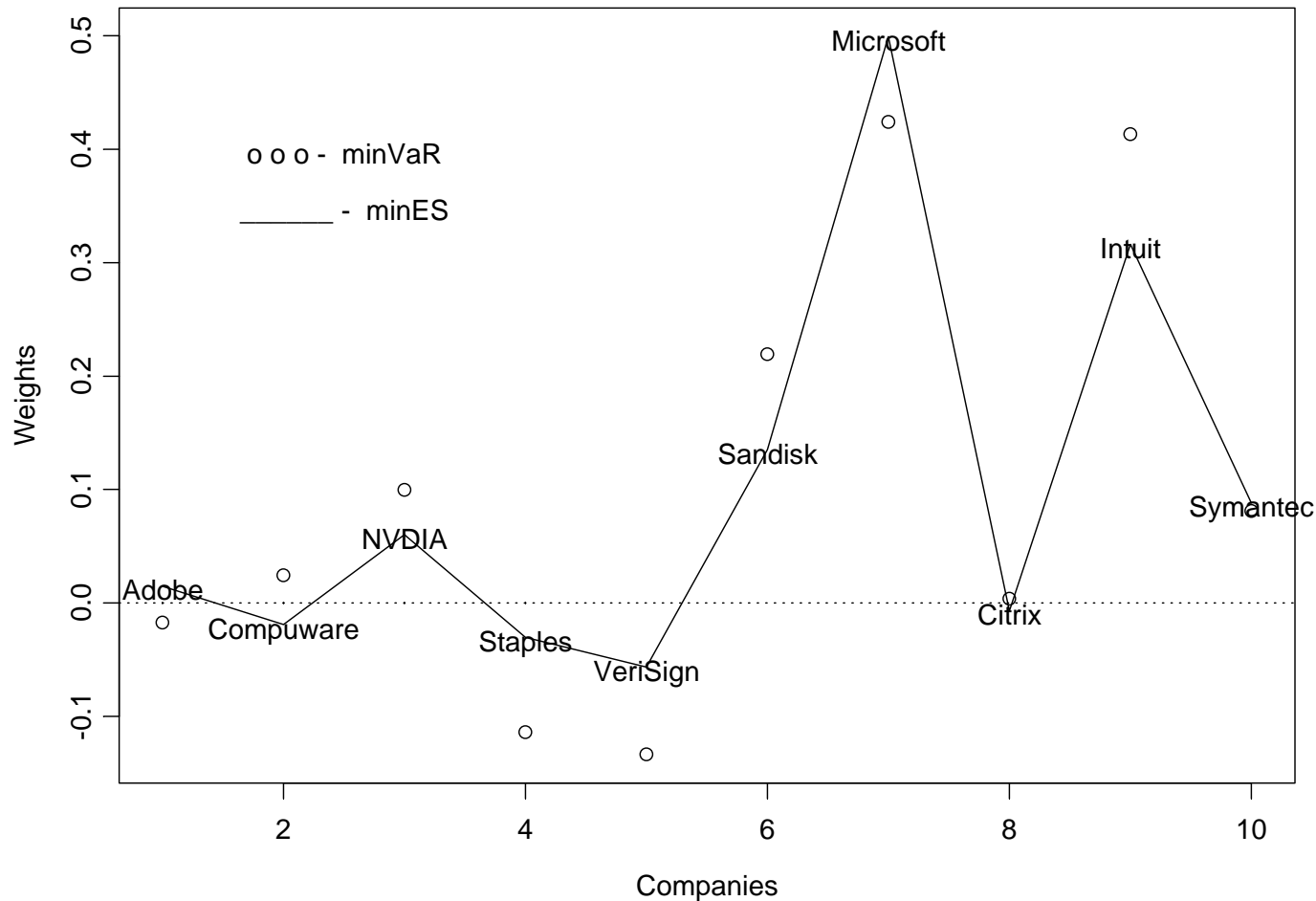
$$ES_q(T) \Rightarrow ES - optimal$$

Figure 1. Lower boundary of VaR and ES q- probability level for GST .



# Closeness between VaR and ES optimal portfolios

Figure 2. VaR and ES optimal portfolios for Norm.



# Distance between portfolios

$$\arg \min_{\mathbf{1}^T \mathbf{x}} VaR_q(L), \arg \min_{\mathbf{1}^T \mathbf{x}} ES_q(L) \in R^n$$

$d =$

$$\sqrt{(\arg \min_{\mathbf{1}^T \mathbf{x}} VaR_q(L) - \arg \min_{\mathbf{1}^T \mathbf{x}} ES_q(L))^T (\arg \min_{\mathbf{1}^T \mathbf{x}} VaR_q(L) - \arg \min_{\mathbf{1}^T \mathbf{x}} ES_q(L))}$$

Denote

$$\delta_q = \lambda_q - z_q = ES_q(Z) - VaR_q(Z)$$

$$C = (\mathbf{1}^T \Sigma^{-1} \mathbf{1})^{-1/2} \sqrt{\Delta^T Q^{-2} \Delta + (\mathbf{1}_1^T Q^{-1} \Delta)^2}$$

**Theorem 4**  $\mathbf{X} \square E_n(\mu, \Sigma, g_n)$ .

$$1) d = [(z_q^2 - B^2)^{-1/2} - (\lambda_q^2 - B^2)^{-1/2}]C \rightarrow 0, q \rightarrow 1$$

$$2) \text{ Let } -z_q^2 \frac{d \log g_1(\frac{1}{2} z_q^2)}{dx} \xrightarrow{q \rightarrow 1} l$$

$$d = \frac{1}{\sqrt{z_q^2 - B^2}} \left[ \frac{1}{l-1} + \frac{1}{2} \left( \alpha_q - \frac{2l-3}{(l-2)^2} \right) + O\left( \left( \alpha - \frac{2l-3}{(l-2)^2} \right)^2 \right) \right] C,$$

$$\alpha_q = \frac{\delta_q (2z_q + \delta_q)}{z_q^2 - B^2} \rightarrow \frac{2l-3}{(l-2)^2}.$$



3) If  $\mathbf{X} \square \mathbf{t}_n(\mu, \Sigma; p, k_{n,p})$ ,

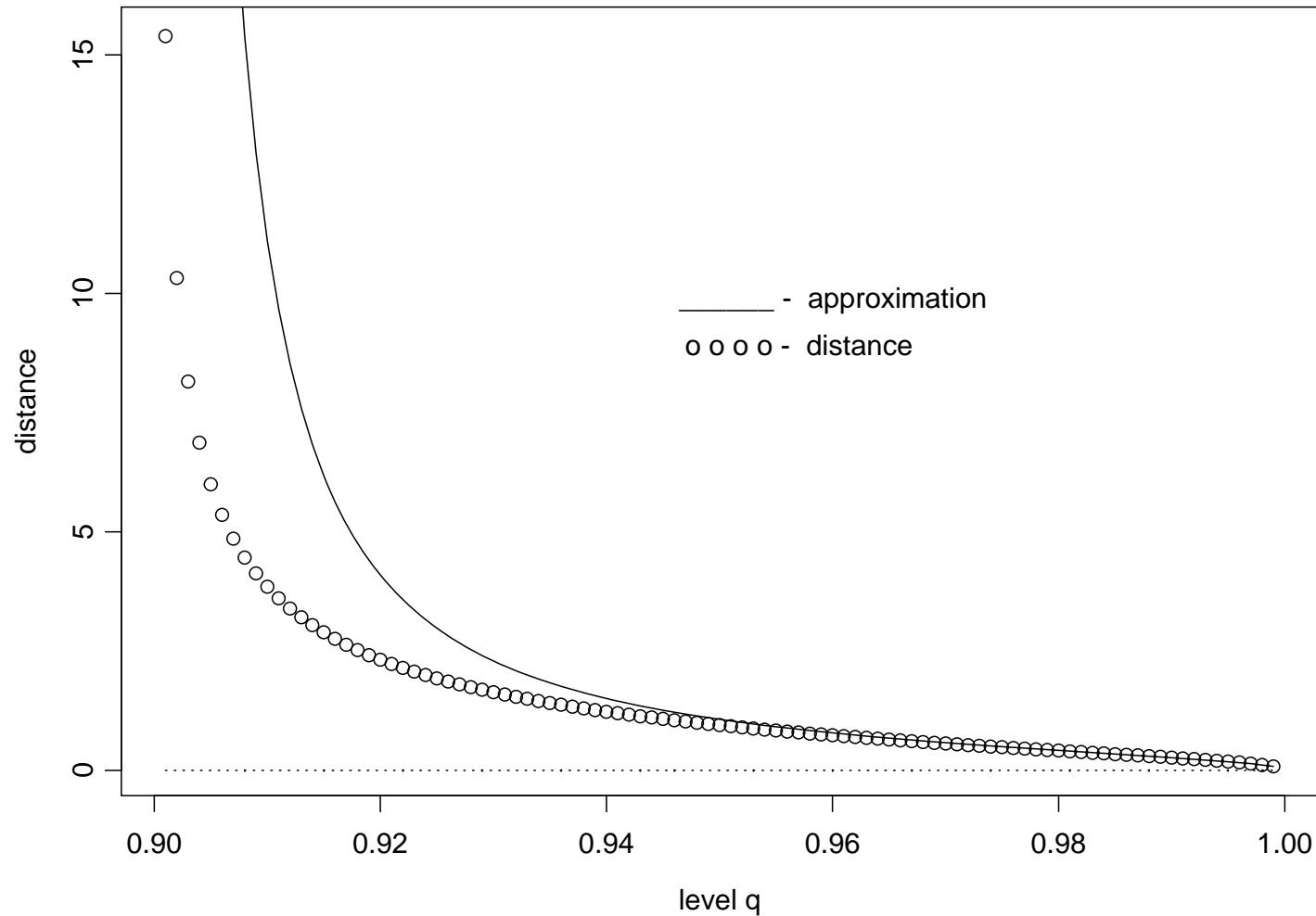
$l = 2p_1$ ,  $p_1 = p - (n-1)/2$ , and

$$d = \frac{1}{\sqrt{z_q^2 - B^2}} \left[ \left( \frac{1}{2p_1 - 1} + \frac{1}{2} \frac{(p_1 - 1)^3}{(p_1 - 1/2)^3} \left( \alpha_q - \frac{p_1 - 3/4}{(p_1 - 1)^2} \right) \right. \right. \\ \left. \left. + O\left( \left( \alpha_q - \frac{p_1 - 3/4}{(p_1 - 1)^2} \right)^2 \right) \right] C.$$

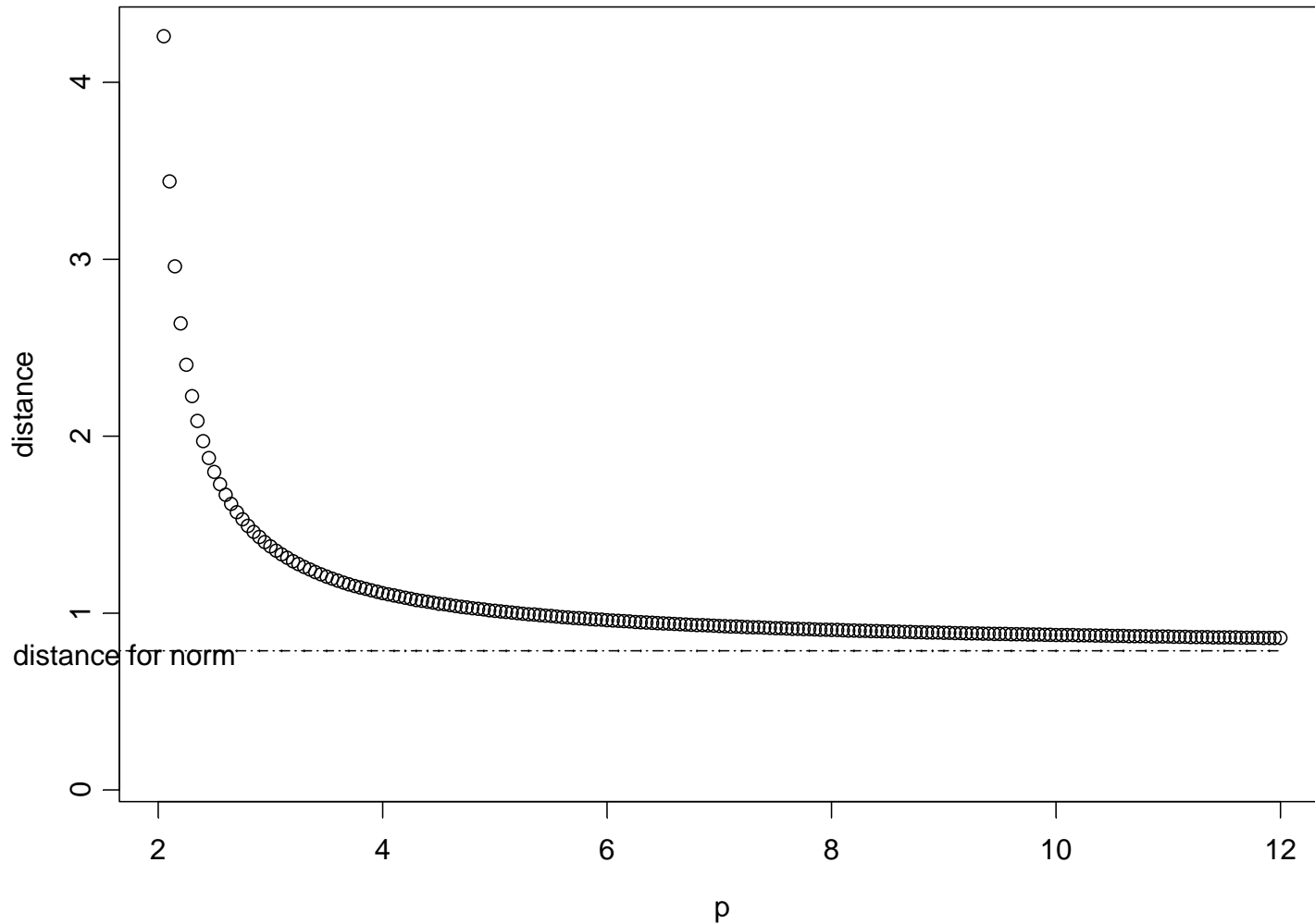
4) If  $\mathbf{X} \square \mathbf{N}_n(\mu, \Sigma)$ ,  $l = \infty$ , and

$$d = \frac{1}{\sqrt{z_q^2 - B^2}} \left[ \frac{1}{2} \alpha_q + O(\alpha_q^2) \right] C$$

# Figure 3. Distance and it's approximation; GST with power parameter $p=1.6$



# Figure 4. Changing distance with increasing of power parameter $p$ of GST



# Conclusion

- Problem of minimization of translation invariant and positive homogeneous risk measures has exact close form solution
- This solution can be easily realized and used for analyzing the influence of parameters of underlying distribution on the portfolio selecting



-The solution is different with that of mean-variance except the case when the portfolio's expected mean is certain

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